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by [Kubo T](#), [Takahashi M](#), [Sato T](#), [Sasaki T](#), [Oka T](#), [Iwasaki K](#)

Affiliation: National Institute of Occupational Safety and Health, 6-21-1, Nagao, Tama-ku, Kawasaki 214-8585, Japan. kubo@h.jniosh.go.jp

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Weekend sleep intervention for workers with habitually short sleep periods

by Tomohide Kubo, PhD,¹ Masaya Takahashi, PhD,¹ Tomoaki Sato, PhD,² Takeshi Sasaki, MSc,¹ Tatsuo Oka, BA,¹ Kenji Iwasaki, PhD¹

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Objective This study was conducted to determine whether extended sleep time during the weekend improves alertness and performance during the subsequent week for workers who are habitually short on sleep time.

Methods Daytime employees in the manufacturing industry [38.3, standard deviation (SD) 8.1 years old, mean weekday sleep ≤ 6 hours] participated in a study that lasted 3 successive weeks. Participants were instructed to stay in bed for ≥ 8 hours between 22:00–09:00 hours on weekends during the first week as a sleep intervention condition and keep their habitual sleep–wake patterns as a habitual weekend sleep condition beginning the weekend of the second week through Thursday of the third week. Half the participants underwent the conditions in one order and the other half in the reverse. Sleep was monitored by an actigraph. A psychomotor vigilance task, subjective fatigue, and blood pressure were measured on Monday and Thursday during the afternoon each week.

Results Sleep duration on weekends was approximately 2 hours longer per day during the intervention. However, sleep duration during weekdays following the intervention returned to shorter periods. Significantly shorter reaction times and a smaller number of lapses on the psychomotor vigilance task were found on Mondays after the intervention than after the habitual weekend sleep. The opposite results, however, were observed on Thursdays.

Conclusions Sleep extension on weekends may be effective in improving alertness and performance during the first days in subsequent weeks among workers with short sleep times. These benefits might be maintained if sufficient sleep duration continues.

Key terms alertness; fatigue; performance; sleepiness; sleep debt; weekend sleep extension.

Working overtime is common in developed countries and has increased in recent years (1). Previous studies have suggested an association between long working hours and adverse health outcomes, such as insufficient sleep, depression, and cardiovascular disease (2–4). In particular, a great deal of research has focused on the influence of overtime work on health problems in Japan (5–7).

Sleep plays an essential role in recovery from work-related fatigue (8). Shortened sleep due to long working hours or overtime may thus increase the likelihood of insufficient recovery from fatigue (9–11) and stress-related diseases (12–15). Therefore, an intervention to ensure adequate sleep is required to protect workers' health and well-being. A potential opportunity for such an intervention is assumed to be available either during weekdays or on weekends. For example, it may be

effective to set a “no-overtime” day in the middle of the week. Alternatively, an intervention on weekends would be more feasible. Indeed, workers tend to extend their sleep duration on weekends to compensate for sleep debt accumulated between Monday–Friday (7).

A potential problem with attempting to sleep longer on weekends involves circadian rhythms. Specifically, delayed wake-up times on weekends have undesirable effects on alertness and performance during the subsequent weekdays (eg, the Blue Monday phenomenon) (16, 17). These disadvantages need to be avoided, even if a certain amount of weekend sleep (ie, sleep compensation) would be important to facilitate recovery from work-induced fatigue. Currently, little is known about which factors (ie, sleep extension or sleep regularity) should be prioritized for promoting fatigue recovery. Given that daytime workers are likely to compensate for

¹ National Institute of Occupational Safety and Health, Kawasaki, Japan.

² Department of Management Design, Faculty of Business Administration, Osaka International University, Osaka, Japan.

Correspondence to: Tomohide Kubo, PhD, National Institute of Occupational Safety and Health, 6-21-1, Nagao, Tama-ku, Kawasaki 214-8585, Japan. [E-mail: kubo@h.jniosh.go.jp]

the sleep debt they accumulate throughout the week by extending their weekend sleep times at the expense of sleep regularity (7), examining the costs and benefits of extended sleep duration during weekends is crucial in terms of occupational health. Unfortunately, to the best of our knowledge, no research has examined the effects of sleep extension during weekends among workers. The present study was conducted to determine whether sleep extension during weekends improves alertness and performance during the subsequent weekdays for workers with habitually short sleep times.

Methods

Participants

Before the study began, a questionnaire was distributed to 402 employees working at an electric company in Kyoto prefecture in order to select possible participants. The respondents were given the morningness–eveningness questionnaire (18) and asked about their habitual sleep time during weekdays and weekends and their health condition (ie, whether they have been sick and whether they take or have taken prescription medications). The selection criteria for participants were set as (i) “neither entirely morning nor evening type,” as evaluated by the questionnaire, (ii) <6 hours mean sleep duration on weekdays, and (iii) taking no prescription medications. Consequently, 49 respondents who met the selection criteria were recruited. Of these, 28 employees [mean 38.8, standard deviation (SD) 8.5 years old, 18 men and 10 women] who were able to attend all sessions of the study were selected as participants. However, two participants withdrew from the study, citing job requirements and unexpectedly poor health conditions, respectively (mean 38.3, SD 8.1 years old, 18 men and 8 women). Regarding their social status, 23 participants cohabited with someone, whereas 2 participants lived alone (1 participant did not answer). Additionally, their weekly working hours in the past month were 35–40 hours for 34.6% (N=9), 41–50 for 42.3% (N=11), 51–60 for 19.2% (N=5), and 61–70 for 3.8% (N=1). Before the study began, the participants were provided with an information sheet outlining the goals of the study and were given the opportunity to ask questions. The Research Ethics Committee of the National Institute of Occupational Safety and Health, Japan, approved this study. All participants gave written informed consent and received payment for their participation.

Study design

This study lasted three successive weeks (from Monday of the first week to Thursday of the third week; see

figure 1). The participants were required to complete two conditions (the intervention and the habitual weekend sleep condition). Half the participants underwent the conditions in one order (eg, from the intervention to the habitual weekend sleep condition) and the other half underwent the conditions in the reverse. They were instructed to stay in bed for ≥ 8 hours between 22:00–09:00 hours on the weekend (Friday, Saturday, and Sunday) of the first week as a sleep intervention and to keep their habitual sleep–wake patterns as a control beginning the weekend of the second week through Thursday of the third week.

Measurements

A battery of tests, which consisted of a psychomotor vigilance task, self-rated fatigue symptoms, and blood pressure, was used to measure the effects of the study conditions.

The psychomotor vigilance task is a sustained-attention reaction time task with a random interstimulus interval of 2–10 seconds (PVT-192, Ambulatory Monitoring, Ardsley, New York, USA) (19). Lapses (reaction times >500 milliseconds) were counted during each 10-minute test as a measure of performance impairment indicative of reduced behavioral alertness. The data were analyzed for speed ($1/\text{reaction time} \times 1000$) and transformed lapses (square root transform [$\text{SQR}(\text{lapses}) + \text{SQR}(\text{lapses} + 1)$]).

Subjective fatigue symptoms were assessed by a questionnaire that asked about work-related feelings of fatigue *Jikaku-sho shirabe* (20), which was proposed by the Industrial Fatigue Research Committee of Japanese Occupational Health in 2002. This questionnaire consists of 25 subjective fatigue symptoms that are categorized into 5 factors of feeling: (i) drowsiness, (ii) instability, (iii) uneasiness, (iv) local pain or dullness, and (v) eyestrain. For each item, participants are requested to estimate the intensity of their feelings using the following scale: “totally disagree,” “agree scarcely,” “agree slightly,” “agree considerably,” and “agree strongly.” These five intensities were assigned scores of 1–5 points, respectively.

The participants’ blood pressure (ES-P2100, TERUMO Corporation, Tokyo, Japan) was measured twice during each test. The mean of these measurements was recorded.

Sleep was monitored by an actigraph (Micro-mini RR, Ambulatory Monitoring, Ardsley, New York, USA) throughout the period of the study. The actigraph was secured on the participants’ non-dominant wrist. Epoch length was set at 1-minute intervals. Sleep data were analyzed by Action-W® version 2.4.20 software (Ambulatory Monitoring, Ardsley, New York, USA) with the Cole-Kripke algorithm (21). In this study, total sleep time

Figure 1. Study design (within-subject design). In the sleep intervention condition, participants were instructed to stay in bed for ≥ 8 hours between 22:00–09:00 hours on weekends (Friday, Saturday, and Sunday) of the 1st week (or 2nd week) as a sleep intervention. Meanwhile, they were required to keep their habitual weekend sleep patterns as a control on the weekend of the 2nd week (or 1st week), followed until Thursday of the 3rd week.

(TST), sleep efficiency, bedtime, and wake time were analyzed to assess the quality of sleep. Sleep efficiency was calculated as: (ratio of TST to total time in bed) $\times$ (100).

Procedure

The present study lasted three weeks. Participants were required to follow their usual sleep–wake cycles in the first week as a baseline. In the intervention condition, they were instructed to remain in bed for ≥ 8 hours between 22:00–09:00 hours and prohibited from taking a daytime nap on weekends (Friday, Saturday, and Sunday). The participants were encouraged to spend the weekends following the instructions as closely as possible. No other assistance was provided (eg, a telephone reminder). In the habitual weekend sleep condition, participants were required to keep their habitual sleep–wake patterns on the weekend.

The test battery was conducted on Monday and Thursday during the afternoon in all three weeks. Participants were instructed to arrive at a testing room between 14:00–15:00 hours. The order of measurements was subjective fatigue (5 minutes), blood pressure (5 minutes), and the psychomotor vigilance task (10 minutes). During the test period, participants were required to switch off their cell phones and take off their wrist-watches so that they could concentrate on the test. They also practiced the psychomotor vigilance task for a few minutes before each test period.

Data analysis

Data for the test battery and actigraph were analyzed using two-way mixed-model analysis of variance (ANOVA). The ANOVA model consisted of subjects as the random effect and the following fixed effects: condition (sleep intervention versus habitual weekend sleep) and day (test battery: 4 levels [Monday, Thursday, Monday, Thursday] and actigraph: 7 levels [Monday, Tuesday, Wednesday, Thursday, Friday, Saturday, Sun-

day]). The main interest was in the interaction effect. Significant interaction was assessed using post-hoc *t*-tests to reveal whether there was a significant difference between conditions. All statistical analyses were performed using SPSS 17.0 for Windows (SPSS Inc, Chicago, IL, USA).

Results

The results of mixed-model ANOVA for all measurements are presented in table 1.

Sleep

For sleep parameters, significant interactions were observed in TST and bedtime ($F_{6,325}=9.74$, $P<0.001$ and $F_{6,325}=17.21$, $P<0.001$, respectively; figure 2). Post-hoc analyses showed that weekend sleep patterns were significantly longer in TST and participants had earlier bedtimes in the intervention than the habitual weekend sleep condition (all $P<0.05$). However, no significant interaction was observed between sleep efficiency and wake time. Regarding the main effect of condition, significant differences were observed in TST ($F_{1,325}=44.56$, $P<0.001$), bedtime ($F_{1,325}=50.47$, $P<0.001$), and wake time ($F_{1,325}=5.84$, $P=0.016$), but not sleep efficiency. The main effect of day was significant for TST ($F_{6,325}=23.64$, $P<0.001$), sleep efficiency ($F_{6,325}=2.93$, $P<0.001$), bedtime ($F_{6,325}=8.10$, $P<0.001$), and wake time ($F_{6,325}=51.62$, $P<0.001$).

Sustained attention

Psychomotor vigilance task performance showed significant interactions in reaction times and lapses ($F_{3,175}=6.01$, $P=0.001$ and $F_{3,175}=4.53$, $P=0.004$, respectively; figure 3). Significantly shorter reaction times ($P<0.05$) and fewer lapses ($P<0.05$) were observed

Table 1. Mixed-model analysis of variance results.

	Mixed-model analysis of variance					
	Condition		Day		Interaction	
	F-value	P-value	F-value	P-value	F-value	P-value
Actigraphic sleep parameters ^a						
Total sleep time	44.56	0.000	23.64	0.000	9.74	0.000
Sleep efficiency	0.82	0.367	2.93	0.008	1.97	0.069
Bedtime	50.47	0.000	8.10	0.000	17.21	0.000
Wake time	5.84	0.016	51.62	0.000	1.39	0.219
Psychomotor vigilance task ^b						
Reaction time	0.02	0.890	0.41	0.745	6.01	0.001
Lapse	0.06	0.803	0.62	0.606	4.53	0.004
Subjective fatigue questionnaire ^b						
Drowsiness	2.90	0.090	4.29	0.006	1.81	0.148
Eyestrain	0.18	0.670	6.14	0.001	0.34	0.794
Instability	0.09	0.759	11.54	0.000	0.18	0.910
Uneasiness	0.24	0.627	7.84	0.000	0.08	0.971
Local pain or dullness	0.00	0.988	2.44	0.066	0.02	0.995
Blood pressure ^b						
Systolic blood pressure	1.21	0.274	13.62	0.000	1.69	0.171
Diastolic blood pressure	0.01	0.906	2.38	0.071	0.24	0.869

^a Condition=df(1,325); Day and Interaction=df(6,325).^b Condition=df(1,175); Day and Interaction=df(3,175).

on Mondays following the intervention condition, as compared with the habitual weekend sleep condition. In contrast, the intervention condition showed significantly longer reaction times and a larger number of lapses on Thursdays than the habitual weekend sleep condition ($P<0.05$). No significant main effects of condition and day were observed.

Subjective fatigue

No significant interactions were found between the intervention and the habitual weekend sleep condition in any of the five subscales of the questionnaire for work-related feelings of fatigue (figure 4). No significant main effects of condition were observed for any of the five subscales. However, the main effects of day were significant for drowsiness ($F_{3,175}=4.29$, $P<0.001$), eyestrain ($F_{3,175}=6.14$, $P=0.001$), instability ($F_{3,175}=11.54$, $P<0.001$), and uneasiness ($F_{3,175}=7.84$, $P<0.001$), but not for local pain or dullness.

Blood pressure

The two parameters of blood pressure showed no significant interactions between the intervention and the habitual weekend sleep condition (figure 5). No main effects of condition were found for the two parameters. The main effects of day were significant for systolic blood pressure ($F_{3,175}=13.62$, $P<0.001$), but not for diastolic blood pressure.

Discussion

The purpose of this study was to examine whether weekend sleep extension contributes to improved alertness and performance during the following week among day-time workers with habitually short sleep duration. TST for the intervention condition was 1–2 hours longer per day than for the habitual weekend sleep condition (from 5–7 hours). However, TST during weekdays after the intervention again returned to 5 hours. Performance on the psychomotor vigilance task significantly improved on Monday after the intervention. Nevertheless, significantly longer reaction times and a larger number of lapses on the psychomotor vigilance task were observed on Thursday after the intervention than after the habitual weekend sleep condition. With respect to subjective fatigue and blood pressure, significant favorable effects of sleep extension were not found. Taken together, the present findings suggest that sleep extension on weekends produced favorable influences on alertness and performance during the first day of the subsequent week, though these benefits dissipated after that.

As expected, an improvement in alertness and performance was observed on the Monday following the sleep intervention. This finding agrees with previous studies, which reported better neurobehavioral performance and less subjective sleepiness by recovery sleep following shortened sleep protocols (22–24). However, these findings were inconsistent with the findings of

Figure 2. Mean and standard error (SE) of actigraphic sleep parameters before, during and after sleep intervention. Shared area indicates the period of the sleep intervention. * Significant difference from control condition at $P < 0.05$.

Taylor et al (16). The study reported that a 2-hour delay in wake-up time on the weekend deteriorated alertness and performance until the middle of the next week. One explanation for this discrepancy is that the participants in the present study were habitually sleep deprived, whereas normal sleepers (those who sleep approximately 8 hours) participated in the previous study. Hence, the effects of sleep extension might have worked to cancel potential sleep debt that participants of the present study had been accumulating. In the previous study (16), extended night-time sleep (mean 9.8 hours) might act as “extra” sleep. Sleep extension is shown to result in less clear improvements in alertness and performance for individuals without sleep debt (23, 25, 26). Another contributing factor

may be wake time. The participants in the present study woke up by 09.00 hours at the latest (figure 2), whereas those in Taylor et al’s study awoke at approximately 10:30 hours. Earlier wake time may prevent circadian phase delay, causing negative influences on daytime function on subsequent weekdays.

An important question in the present study is how long the benefit of weekend sleep extension lasted following the subsequent Monday. It was unexpected that a significantly longer reaction time and a larger number of lapses on the psychomotor vigilance task were found on the Thursday after the intervention than after the habitual weekend sleep condition (figure 3). This finding might be related to the immediate return of sleep

Figure 3. Mean and standard error (SE) of psychomotor vigilance task for reaction times and lapses before and after the sleep intervention. Shared area indicates the period of the sleep intervention. * Significant difference from control condition at $P<0.05$.

Figure 4. Mean and standard error (SE) of subjective fatigue questionnaire for drowsiness and eyestrain before and after the sleep intervention. Shared area indicates the period of the sleep intervention.

Figure 5. Mean and standard error (SE) of blood pressure before and after the sleep intervention. Shared area indicates the period of the sleep intervention.

duration following the intervention (from 7 to 5 hours). In the present study, sleep duration on the weekends was approximately 2 hours longer per day for the intervention than the habitual weekend sleep condition (figure 2). However, sleep duration on weekdays after the intervention returned to 5 hours. Thus, we assume that the immediate return from the longer weekend sleep duration to the shorter weekday sleep duration may cause impaired neurobehavioral performance in the intervention condition. Additionally, given that the cumulative effects of mild sleep restriction became evident within a few days after the start of the restriction (23, 24, 27), the favorable influences of sleep extension in this study may be cancelled due to cumulative effects.

However, a recent study indicated that one week of sleep extension not only improved performance and alertness during the subsequent period of sleep restriction but also facilitated recovery in the post-restriction period (22). These findings were inconsistent with the impairments of neurobehavioral performance on the Thursday after the intervention. In the previous research, the targeted participants were normal sleepers, but in the present research, short sleepers. Namely, the effects of sleep extension in this study might work as compensation for accumulated sleep debt among short sleepers. In contrast, one week of extended sleep times might result in the benefits of sleep extension for normal sleepers. When sleep duration is adequate not only on weekends but also weekdays, the effects of sleep extension can be maintained.

Furthermore, several possibilities exist regarding why the impaired neurobehavioral performance occurred on Thursday after the intervention. Responses to the intervention condition varied among the participants. In short, the standard deviation of reaction times and lapses was 2 or 3 times larger on Thursday after the intervention than on Monday after the intervention (SD for Monday and Thursday varied from 26.1–52.1 in reaction times and 0.9–3.1 in lapses). Therefore, some of the workers showed greater decreases in the performance of a reaction time task toward Thursday, suggesting inter-individual differences in response to a manipulated sleep–wake pattern (28, 29). On the contrary, the performance data in the habitual weekend sleep condition indicate that deteriorated performance on the subsequent Monday was restored to the baseline level on Thursday. Consequently, the combined effects of these two factors are likely to lead to the current result of impaired neurobehavioral performance on Thursday after the intervention. Alternatively, the workload intensity would be unexpectedly changed on Thursday after the intervention. Another reason might be that a type-1 error because the sample size in this study was not large. Taken together, three days (Friday to Sunday) of sleep extension on weekends may have sustained the benefits for only a few days following the intervention, as far as sufficient sleep duration on the subsequent weekdays could be not ensured.

This study had some limitations. First, participants in the present study slept <6 hours on weekdays, despite the fact that they did not work extremely long hours

(about 80% of workers in this study had mean weekly working hours of <50 hours). It is possible that they voluntarily curtailed their sleep duration. We interviewed some of the participants about the causes of their habitually short sleep times, but no clear responses emerged. In addition, the participating workers might be categorized as individuals with behaviorally induced insufficient sleep syndrome. Such a possibility, however, would be quite low, given the clinical presentation and the diagnostic criteria for this syndrome (30).

Second, it remains unclear whether the weekend sleep extension had positive effects on work productivity because the participants' work output was not evaluated.

Third, we did not measure how participants spent their leisure time. Recently, a great deal of research has been conducted on the importance of leisure activity in fatigue recovery (31–33). Moreover, findings from recent studies indicate that increased control of working time and/or days off plays a crucial role in improving workers' health (34). Therefore, the benefits of weekend sleep extension are likely to vary depending on the types of leisure activity in which the participants engage (eg, active or passive leisure) and/or working style.

Fourth, the measurements in this study could have been more optimal through the selection of different parameters, such as a circadian phase marker (eg, melatonin). Besides, increased frequency of testing, rather than only on Monday and Thursday, might further clarify the possible effects of the intervention. Furthermore, since sleep data in this study were measured by actigraphy instead of polysomnography, sleep might have been overestimated (35). Thus, it is possible that sleep extension on weekends was <2 hours per day.

Fifth, the present study did not address the long-term effects of sleep extension on weekends, and it remains unclear what they would be. But we speculate that the weekend sleep extension may yield favorable effects only on a few days of the following week because sleep duration on weekdays would be short, even if the intervention were repeated for a long period. Both organizational and individual efforts should thus be directed to ensuring sufficient duration of sleep not only on the weekends but also on weekdays.

Finally, it may be difficult to continue the weekend sleep intervention over a long period, given the present data showing an immediate return to short sleep following the weekends. We believe that the participants' high compliance was achieved because of the single short-term (2–3 days) intervention that spanned a weekend.

The sleep intervention in this study succeeded in lengthening the sleep duration of workers with habitually short sleep periods. Longer weekend sleep resulted in acute improvements of alertness and performance, but the improvements were not present later during the

workweek. However, the benefits of weekend sleep extension might be maintained if the participants receive sufficient sleep during the post-intervention week. Our present findings suggest the significance of vouchsafing adequate sleep nightly for optimal levels of worker safety and health.

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